# Lecture 15-16 Pose Estimation – Gaussian Process

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### MS KINECT

http://www.youtube.com/watch?v=p2qlHoxPioM

The KINECT body pose estimation is achieved by randomised regression forest techniques.

We learn pose estimation as a regression problem, and Gaussian process as a cutting edge regression method.

We see it through the case study (slide credits to): Semi-supervised Multi-valued Regression,

Navaratnam, Fitzgibbon, Cipolla, ICCV07,

where practical challenges addressed are 1)multi-valued regression and 2)sparsity of data.





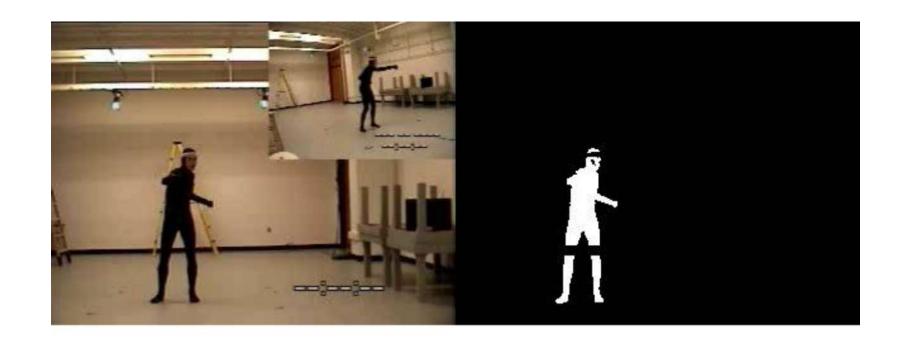


Image I

Pose  $\theta$ 

e.g. Urtasun, Fleet, Hertzmann, Fua; ICCV 2005.

A mapping function is learnt from the input image I to the pose vector  $\theta$ , which is taken as a continuous variable.

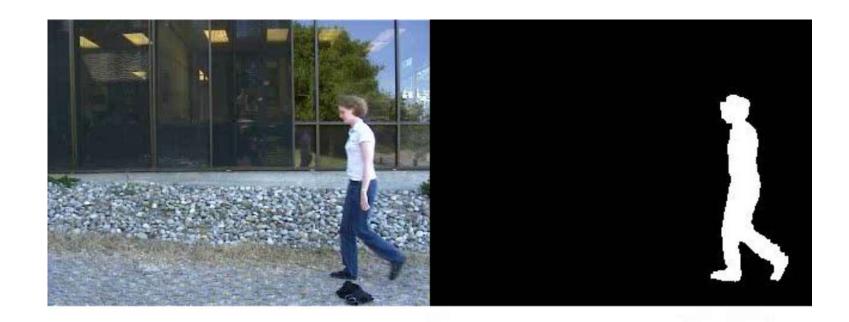






Given an image (top left), e.g. a silhouette (top right) is obtained by background subtraction techniques (<a href="http://en.wikipedia.org/wiki/Background subtraction">http://en.wikipedia.org/wiki/Background subtraction</a>). The estimated 3d pose is shown at two camera angels (bottom left and right).

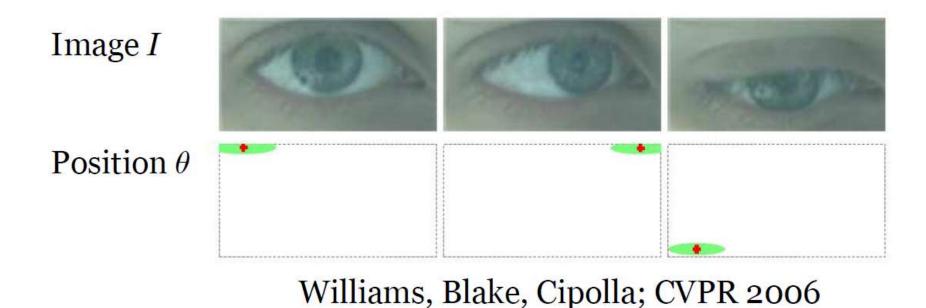
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We attempt to map 2D image space to 3D pose space. There is inherent ambiguity in pose estimation (as an example in the above).



Eye tracking can be tackled as a regression problem, where the input is an image I and the output is a eye location.

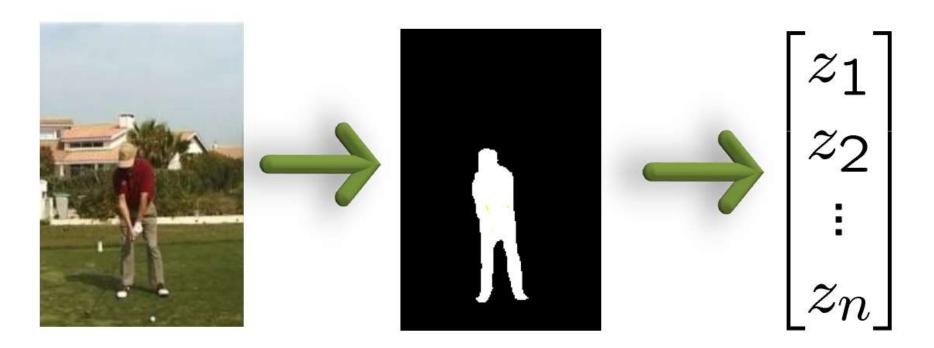


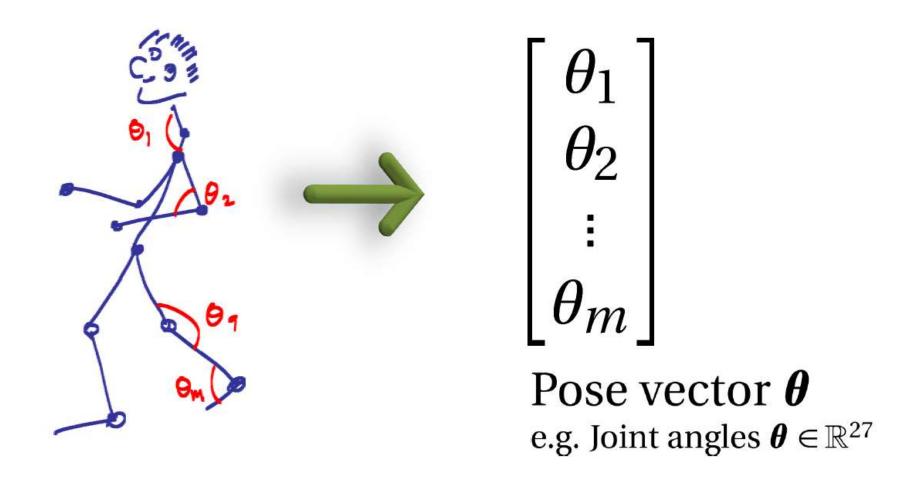
Image I

Feature vector  $\mathbf{z}$  e.g. Shape contexts on silhouette,  $\mathbf{z} \in \mathbb{R}^{40}$ 

Typical image processing steps:

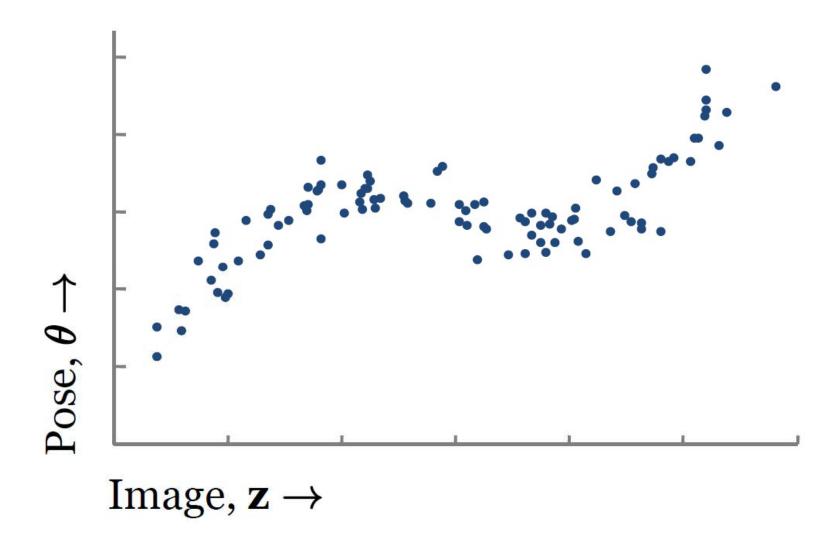
Given an image, a silhouette is segmented.

A shape descriptor is applied to the silhouette to yield a finite dimensional vector. (Belongie and Malik, Matching with Shape Contexts, 2000)

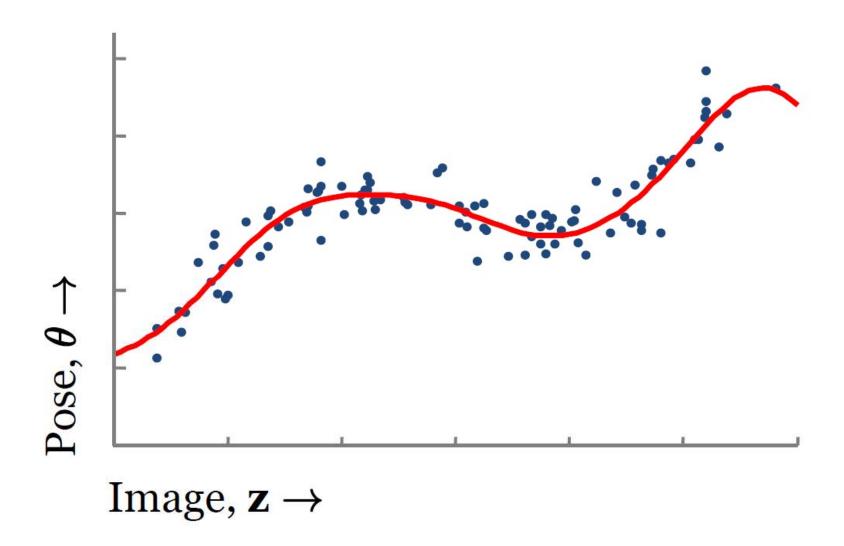


The output is a vector of m joint angles.

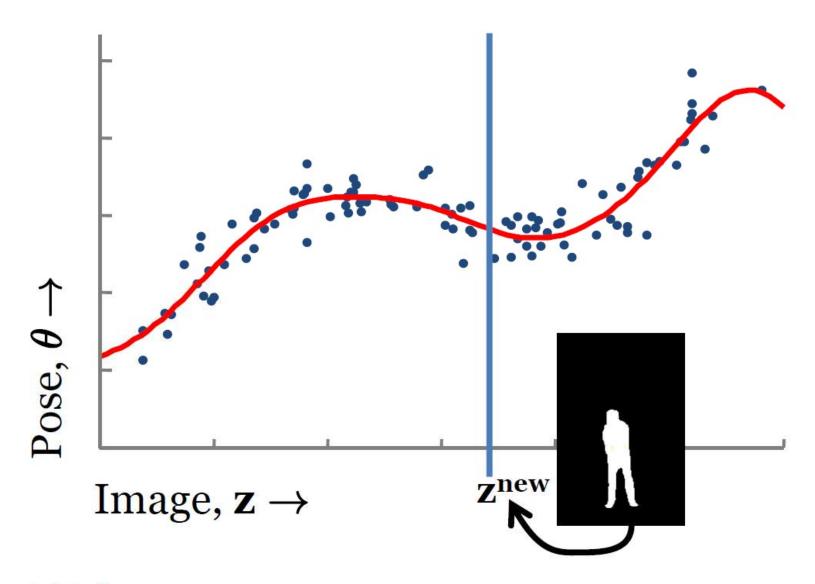
## 1. Obtain training samples $(\mathbf{z}_1, \boldsymbol{\theta}_1)...(\mathbf{z}_N, \boldsymbol{\theta}_N)$



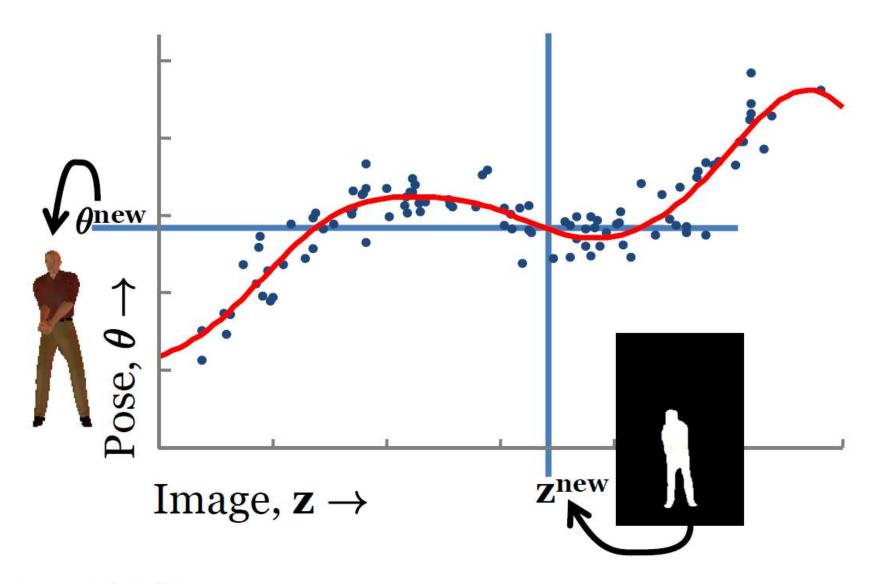
### 2. Training: Fit function $\theta = f(\mathbf{z})$ .



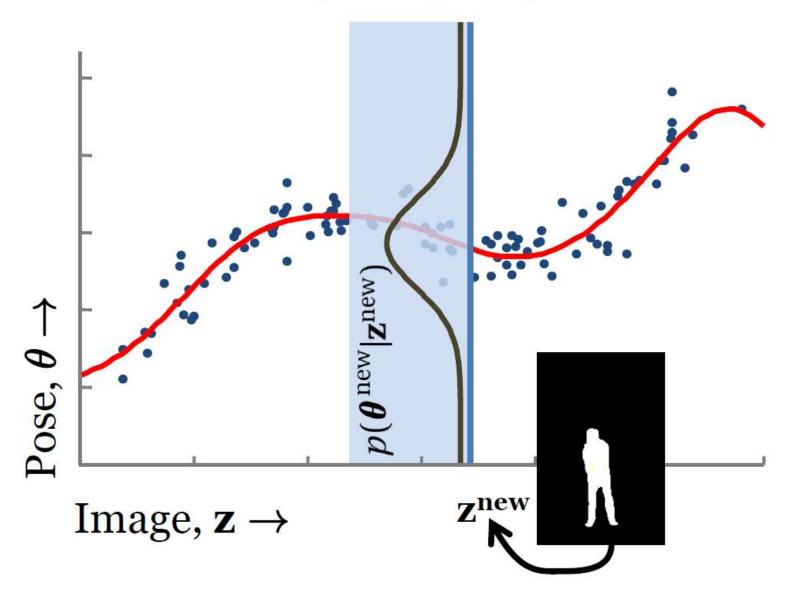
3. Given new image,  $\mathbf{z}^{\text{new}}$ , compute  $\boldsymbol{\theta}^{\text{new}} = f(\mathbf{z}^{\text{new}})$ .



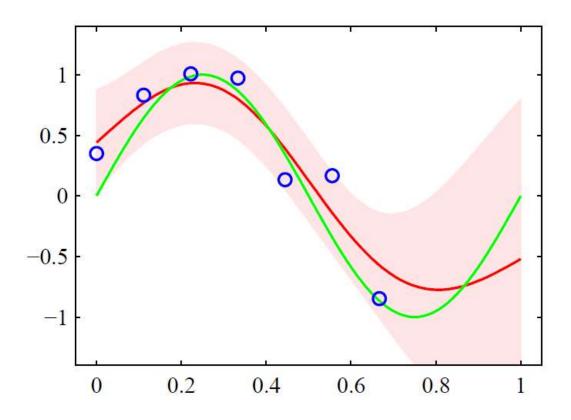
# 3. Given new image, $\mathbf{z}^{\text{new}}$ , compute $\boldsymbol{\theta}^{\text{new}} = f(\mathbf{z}^{\text{new}})$ .



# 3. Or, more usefully, compute $p(\boldsymbol{\theta}^{\text{new}}|\mathbf{z}^{\text{new}})$ .

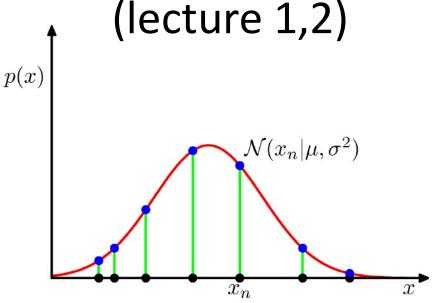


## **Gaussian Process**



## Imperial College

# London Review of Gaussian density estimation (lecture 1.2)



i.i.d.(independent identical distributed)

$$\mathbf{x} = (x_1, ..., x_N)^T$$

We want to find the Gaussian parameters from the given data.

The problem is to find the parameters by maximising the posterior probability

$$p(\mu, \sigma^2 | \mathbf{x}) = \frac{p(\mathbf{x} | \mu, \sigma^2) p(\mu, \sigma^2)}{p(\mathbf{x})}$$

# London Review of Gaussian density estimation (lecture 1,2)

$$p(\mu, \sigma^2 | \mathbf{x}) = \frac{p(\mathbf{x} | \mu, \sigma^2) p(\mu, \sigma^2)}{p(\mathbf{x})}$$

We do not have priors on the parameters and data, thus we maximise the (log) likelihood function instead.

$$p(\mathbf{x}|\mu,\sigma^2) = \prod_{n=1}^{N} \mathcal{N}(x_n|\mu,\sigma^2).$$

$$\ln p(\mathbf{x}|\mu,\sigma^2) = -\frac{1}{2\sigma^2} \sum_{n=1}^{N} (x_n - \mu)^2 - \frac{N}{2} \ln \sigma^2 - \frac{N}{2} \ln(2\pi).$$

Maximum Likelihood (ML) vs Maximum A Posterior (MAP)

solutions:

$$P(X|Y)$$

$$P(X|Y)P(Y)$$
(e.g. Gaussian Process)

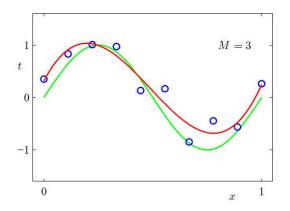
# Imperial College London Review of polynomial curve fitting (lecture 1,2)

We want to fit a polynomial curve to given data pairs (x,t).

$$y(x, \mathbf{w}) = w_0 + w_1 x + w_2 x^2 + \dots + w_M x^M = \sum_{j=0}^M w_j x^j = \mathbf{w}^T \mathbf{x}$$
where  $\mathbf{w} = \begin{bmatrix} w_0 \\ w_1 \\ \vdots \end{bmatrix} \quad \mathbf{x} = \begin{bmatrix} 1 \\ x \\ x^2 \\ \vdots \end{bmatrix}$ 

The objective ftn to minimise is  $E(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^{N} \{y(x_n, \mathbf{w}) - t_n\}^2 + \frac{\lambda}{2} ||\mathbf{w}||^2$ 

where 
$$||\mathbf{w}||^2 = \mathbf{w}^T \mathbf{w} = w_0^2 + w_1^2 ... + w_M^2$$



### Gaussian Processes

- Gaussian Processes: extends the role of kernels to probabilistic discriminative models. We shall see how kernels arise in a Bayesian setting.
- We consider linear regression example and derive the predictive distribution by working in terms of distributions over functions  $y(\mathbf{x}, \mathbf{w})$ .

$$y(\mathbf{x}) = \mathbf{w}^T \phi(\mathbf{x})$$

Consider a prior distribution over w by an isotropic Gaussian of the form

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{w}|\mathbf{0}, \alpha^{-1}\mathbf{I})$$

where the hyperparameter  $\alpha$  is the precision (inverse variance).

The probability distribution over w thus induces

- $\longrightarrow$  a probability distribution over functions  $y(\mathbf{x})$ .
- From given a data set  $\mathbf{x}_1, ..., \mathbf{x}_N$ , we want to know the joint distribution of the function values  $y(\mathbf{x}_1), ..., y(\mathbf{x}_N)$ . We denote the vector  $\mathbf{y}$  that has elements  $y_n = y(\mathbf{x}_n)$ . Then we have

$$\mathbf{y} = \Phi \mathbf{w}$$

where  $\Phi$  is the matrix with elements  $\Phi_{nk} = \phi_k(\mathbf{x}_n)$ .

Note y is a linear combination of Gaussian distributed variables given by the elements of w and hence is itself Gaussian. It has its mean and covariance as follows.

$$\mathbb{E}[\mathbf{y}] = \Phi \mathbb{E}[\mathbf{w}] = \mathbf{0} \qquad \mathbf{y} \sim N(\mathbf{y}|\mathbf{0}, \mathbf{K})$$

$$\operatorname{cov}[\mathbf{y}] = \mathbb{E}[\mathbf{y}\mathbf{y}^T] = \Phi \mathbb{E}[\mathbf{w}\mathbf{w}^T]\Phi^T = \frac{1}{\alpha}\Phi\Phi^T = \mathbf{K}$$

where **K** is the Gram matrix that has

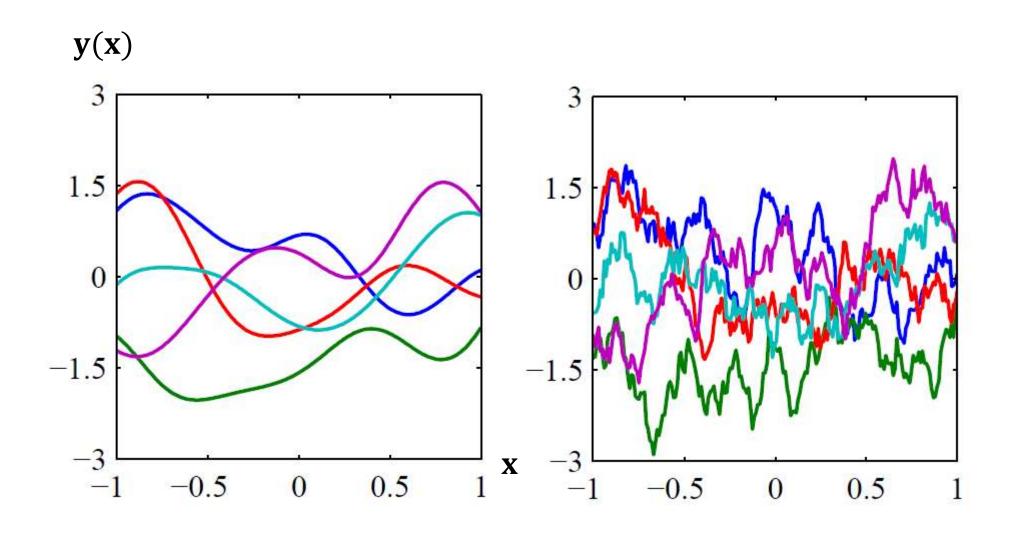
$$K_{nm} = k(\mathbf{x}_n, \mathbf{x}_m) = \frac{1}{\alpha} \phi(\mathbf{x}_n)^T \phi(\mathbf{x}_m).$$

• The figure below shows samples of functions drawn from Gaussian processes for the Gaussian kernel

$$k(x, x') = \exp(-||x - x'||^2/2\sigma^2)$$

(left) and the exponential kernel (right) given by

$$k(x, x') = \exp(-\theta |x - x'|)$$



### Imperial College

# Gaussian Processes for Regression

We consider the noise on the observed target values by

$$t_n = y_n + \epsilon_n$$

where  $y_n = y(\mathbf{x}_n)$ , and  $\epsilon_n$  is a random noise variable.

We consider noise processes that have a Gaussian distribution, so that

$$p(t_n|y_n) = \mathcal{N}(t_n|y_n, \beta^{-1})$$

where  $\beta$  is a hyperparameter representing the precision of the noise.

• The noise is independent for each data point. The joint distribution of the target values  $\mathbf{t} = (t_1, ..., t_N)^T$  conditioned on  $\mathbf{y} = (y_1, ..., y_N)^T$  is an isotropic Gaussian of the form

$$p(\mathbf{t}|\mathbf{y}) = \mathcal{N}(\mathbf{t}|\mathbf{y}, \beta^{-1}\mathbf{I}_N)$$

From the Gaussian process, the marginal distribution p(y) is a Gaussian whose mean is zero and covariance is a Gram matrix K i.e.

$$p(\mathbf{y}) = \mathcal{N}(\mathbf{y}|\mathbf{0}, \mathbf{K}).$$

• To find the marginal distribution  $p(\mathbf{t})$  for given input values, we make use of the following (proof skipped).

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x}|\mu, \Lambda^{-1})$$
$$p(\mathbf{y}|\mathbf{x}) = \mathcal{N}(\mathbf{y}|\mathbf{A}\mathbf{x} + \mathbf{b}, \mathbf{L}^{-1}),$$

the marginal distribution becomes

$$p(\mathbf{y}) = \mathcal{N}(\mathbf{y}|\mathbf{A}\mu + \mathbf{b}, \mathbf{L}^{-1} + \mathbf{A}\Lambda^{-1}\mathbf{A}^T)$$

Thus, we have

$$p(\mathbf{t}) = \int p(\mathbf{t}|\mathbf{y})p(\mathbf{y})d\mathbf{y} = \mathcal{N}(\mathbf{t}|\mathbf{0}, \mathbf{C})$$

where

$$C(\mathbf{x}_n, \mathbf{x}_m) = k(\mathbf{x}_n, \mathbf{x}_m) + \beta^{-1}\delta_{nm}.$$

Their covariances simply add, as the two Gaussian sources of randomness i.e.  $y(\mathbf{x})$ ,  $\epsilon$  are independent.

• One widely used kernel function for Gaussian process regression is

$$k(\mathbf{x}_n, \mathbf{x}_m) = \theta_0 \exp\left\{-\frac{\theta_1}{2}||\mathbf{x}_n - \mathbf{x}_m||^2\right\} + \theta_2 + \theta_3 \mathbf{x}_n^T \mathbf{x}_m.$$

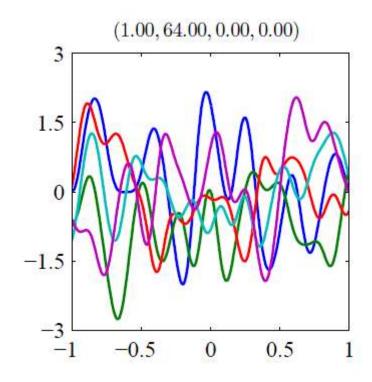


Figure 1: Samples from a Gaussian process prior defined by the kernel function with  $(\theta_0, \theta_1, \theta_2, \theta_3)$ .

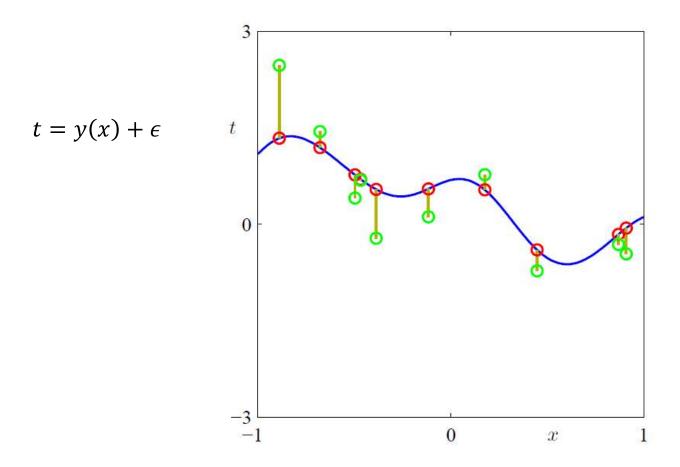


Figure 2: The blue curve shows a sample function from Gaussian process prior over functions, and the red points show the values of  $y_n$ . The greens are the values of  $t_n$  by adding Gaussian noise.

• Our goal is to predict the target value  $t_{N+1}$  for a new input  $\mathbf{x}_{N+1}$ , given a set of training data  $\mathbf{x}_1, ..., \mathbf{x}_N$  and  $\mathbf{t}_N = (t_1, ..., t_N)^T$ .

This requires the evaluation of the predictive distribution  $p(t_{N+1}|\mathbf{t}_N)$ , where we omit the data vectors for notational simplicity.

• The joint distribution is given by

$$p(\mathbf{t}_{N+1}) = \mathcal{N}(\mathbf{t}_{N+1}|\mathbf{0}, \mathbf{C}_{N+1})$$

where  $C_{N+1}$  is an  $(N+1) \times (N+1)$  covariance matrix. We partition the covariance matrix  $as_{N \times N-N \times 1}$ 

$$\mathbf{C}_{N+1} = \left(\begin{array}{cc} \mathbf{C}_N & \mathbf{k} \\ \mathbf{k}^T & c \end{array}\right)$$

where the vector  $\mathbf{k}$  has elements  $k(\mathbf{x}_n, \mathbf{x}_{N+1})$  for n = 1, ..., N, and the scalar  $c = k(\mathbf{x}_{N+1}, \mathbf{x}_{N+1}) + \beta^{-1}$ .

• We use the followings to get the conditional distribution  $p(t_{N+1}|\mathbf{t})$ .

Given

$$\mathbf{x} = \begin{bmatrix} x_a \\ x_b \end{bmatrix}, \quad \mu = \begin{bmatrix} \mu_a \\ \mu_b \end{bmatrix}, \quad \Sigma = \begin{bmatrix} \Sigma_{aa} & \Sigma_{ab} \\ \Sigma_{ba} & \Sigma_{bb} \end{bmatrix}$$

, we have

$$\mu_{a|b} = \mu_a + \Sigma_{ab} \Sigma_{bb}^{-1} (x_b - \mu_b)$$
$$\Sigma_{a|b} = \Sigma_{aa} - \Sigma_{ab} \Sigma_{bb}^{-1} \Sigma_{ba}$$

Therefore, we have  $p(t_{N+1}|\mathbf{t})$  a Gaussian distribution with mean and covariance given by

$$m(\mathbf{x}_{N+1}) = \mathbf{k}^T \mathbf{C}_N^{-1} \mathbf{t}$$

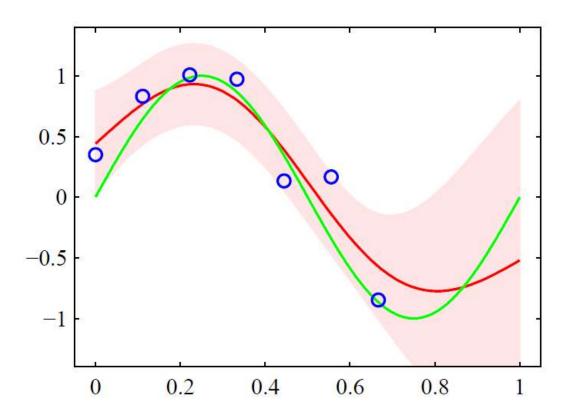
$$\sigma^2(\mathbf{x}_{N+1}) = c - \mathbf{k}^T \mathbf{C}_N^{-1} \mathbf{k}.$$

• The mean of the predictive distribution can be written as a function of  $\mathbf{x}_{N+1}$  as

$$m(\mathbf{x}_{N+1}) = \sum_{n=1}^{N} a_n k(\mathbf{x}_n, \mathbf{x}_{N+1})$$

where  $a_n$  is the *n*-th component of  $\mathbf{C}_N^{-1}\mathbf{t}$ .

These are the key results of Gaussian process regression. An example of Gaussian process regression is shown below.



### Gaussian Process Matlab Toolbox

```
http://www.lce.hut.fi/research/mm/gpstuff/install.shtml
(try demo_regression_robust.m, demo_regression1.m)
```

# Learning Hyperparameters

- The predictions of a Gaussian process depends on the choice of covariance i.e. kernel function. Techniques for learning the hyperparameters are based on the evaluation of the likelihood function  $p(\mathbf{t}|\theta)$ .
- The log likelihood function is given using the standard from of a multivariate Gaussian distribution as

$$\ln p(\mathbf{t}|\theta) = -\frac{1}{2} \ln |\mathbf{C}_N| - \frac{1}{2} \mathbf{t}^T \mathbf{C}_N^{-1} \mathbf{t} - \frac{N}{2} \ln(2\pi).$$

• Maximisation of the log likelihood can be done using the gradient-based optimisation. We can take the derivative w.r.t.  $\theta_i$  as

$$\frac{\partial}{\partial \theta_i} \ln p(\mathbf{t}|\theta) = -\frac{1}{2} \text{Tr} \left( \mathbf{C}_N^{-1} \frac{\partial \mathbf{C}_N}{\partial \theta_i} \right) + \frac{1}{2} \mathbf{t}^T \mathbf{C}_N^{-1} \frac{\partial \mathbf{C}_N}{\partial \theta_i} \mathbf{C}_N^{-1} \mathbf{t}$$

which is obtained by  $\frac{\partial}{\partial x}(\mathbf{A}^{-1}) = -\mathbf{A}^{-1}\frac{\partial \mathbf{A}}{\partial x}\mathbf{A}^{-1}$  and  $\frac{\partial}{\partial x}\ln|\mathbf{A}| = \operatorname{Tr}\left(\mathbf{A}^{-1}\frac{\partial \mathbf{A}}{\partial x}\right)$  (proofs skipped).

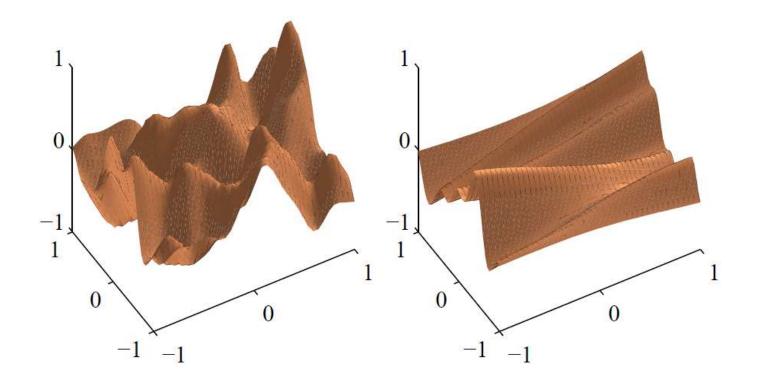
• Because  $\ln p(\mathbf{t}|\theta)$  is in general a nonconvex function, thus having multiple maxima.

### **Automatic Relevance Determination**

• Consider a Gaussian process with a two-dimensional input  $\mathbf{x} = (x_1, x_2)$ , with a kernel function

$$k(\mathbf{x}, \mathbf{x}') = \theta_0 \exp \left\{ -\frac{1}{2} \sum_{i=1}^2 \eta_i (x_i - x_i')^2 \right\}$$

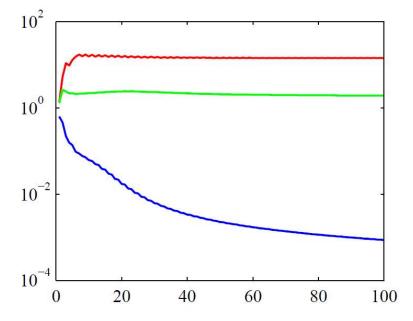
• Samples from the prior over functions  $y(\mathbf{x})$  are shown for two different settings of the precision parameters  $\eta_i$ . As a particular parameter  $\eta_i$  becomes small, the function becomes relatively insensitive to the corresponding input variable  $x_i$ .



• The target variable t is generated by sampling 100 values of  $x_1$  from a Gaussian, evaluating the function  $\sin(2\pi x_1)$ , and then adding Gaussian noise. Values of  $x_2$  are given by copying the values of  $x_1$  and adding noise, and values of  $x_3$  are sampled from an independent Gaussian distribution.

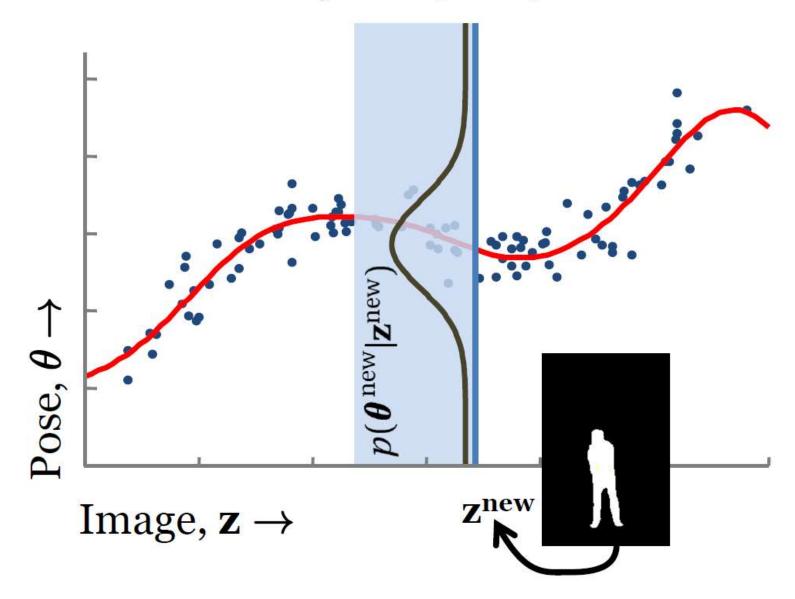
The marginal likelihood is opimised w.r.t.  $\eta_1, \eta_2, \eta_3$  by the gradient algorithm. The figure shows that  $x_1$  is a good predictor of t,  $x_2$  is a more noisy predictor, and  $x_3$  is irrelevant for predicting

t.



# Back to the pose estimation problem

# 3. Or, more usefully, compute $p(\boldsymbol{\theta}^{\text{new}}|\mathbf{z}^{\text{new}})$ .



• It'll never work...

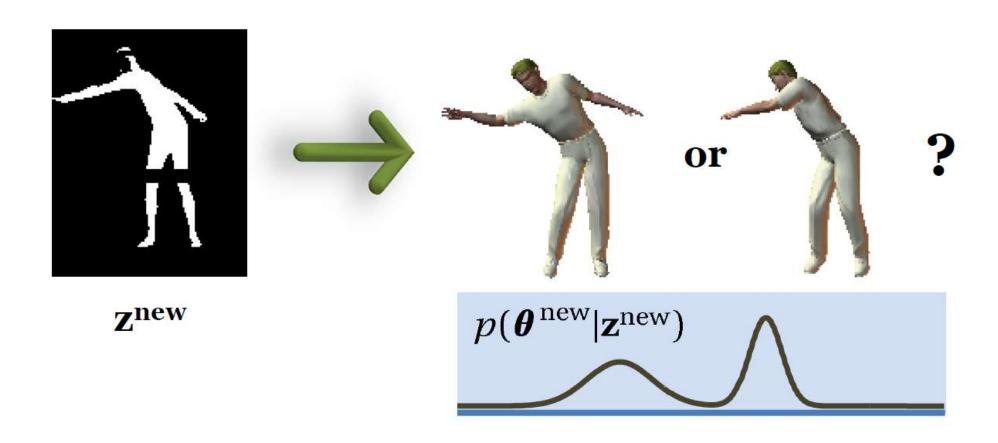
- f is multivalued

- **Z** and  $\theta$  live in high dimensions

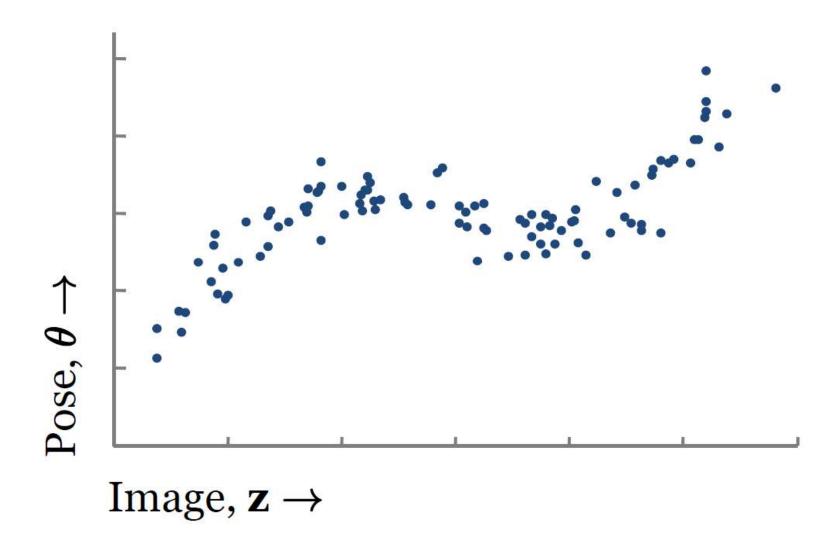
# Multivalued *f*:



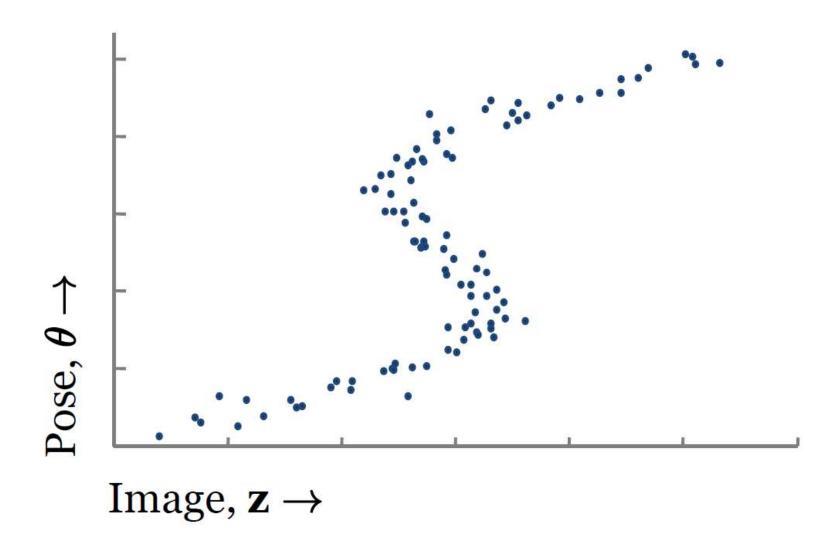
# Multivalued *f*:



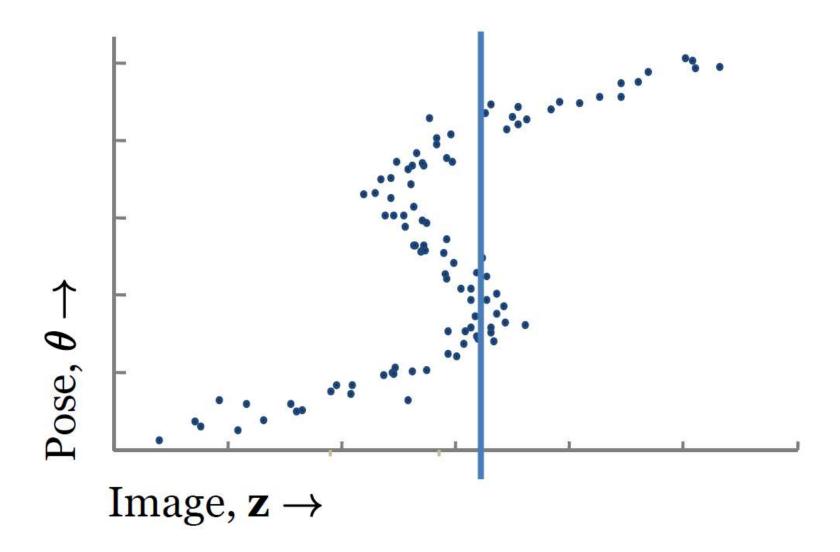
#### Instead of this:



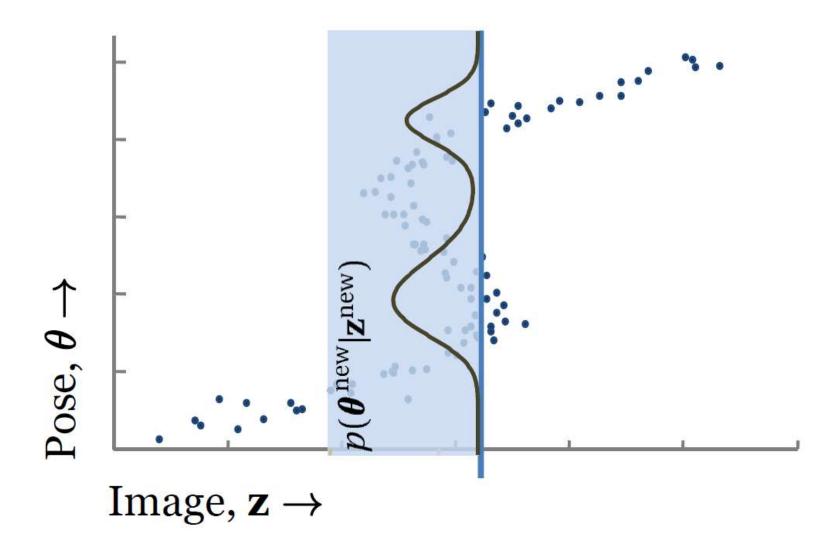






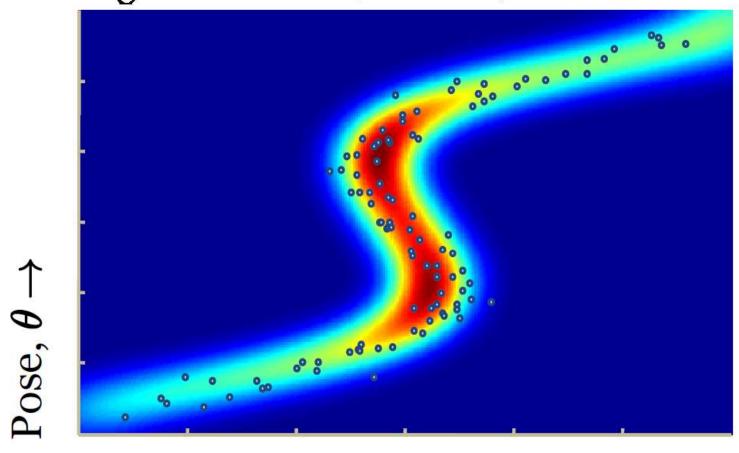








# Instead of fitting $p(\theta | \mathbf{z})$ , fit $p(\theta, \mathbf{z})$ , e.g. with GMM, Parzen, or GPLVM.



Image,  $\mathbf{z} \rightarrow$ 

**GMM**: Gaussian Mixture Model

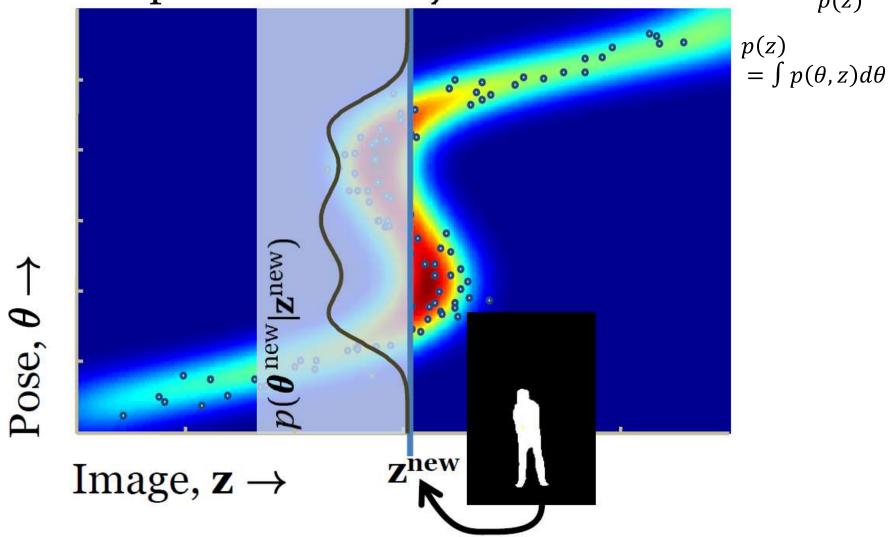
**GPLVM**: Gaussian Process Latent Variable Model



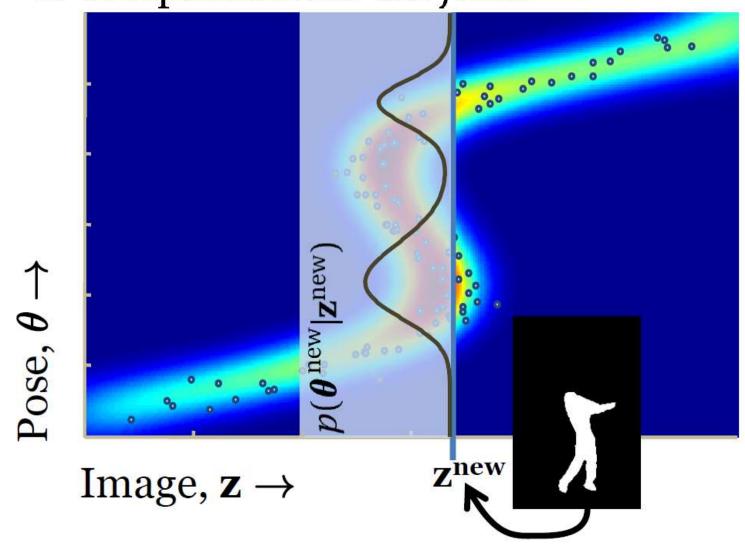
Given new image  $\mathbf{z}^{\text{new}}$ , conditional  $p(\boldsymbol{\theta}|\mathbf{z}^{\text{new}})$ 

is computed from the joint.  $\leftarrow p(\theta|z) = \frac{p(\theta,z)}{p(z)}$ 

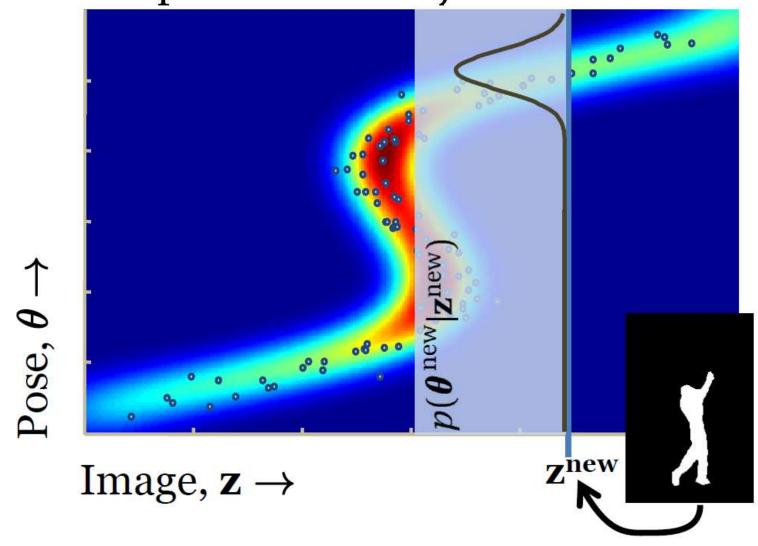
$$p(\theta|z) = \frac{p(\theta, z)}{p(z)}$$

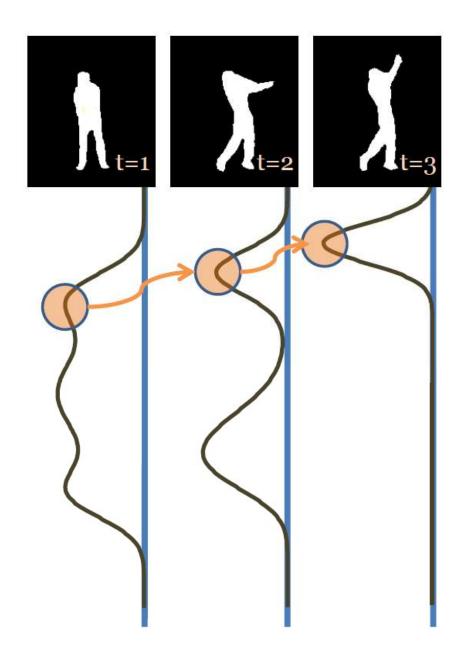


Given new image  $\mathbf{z}^{\text{new}}$ , conditional  $p(\boldsymbol{\theta} | \mathbf{z}^{\text{new}})$  is computed from the joint.



Given new image  $\mathbf{z}^{\text{new}}$ , conditional  $p(\boldsymbol{\theta} | \mathbf{z}^{\text{new}})$  is computed from the joint.



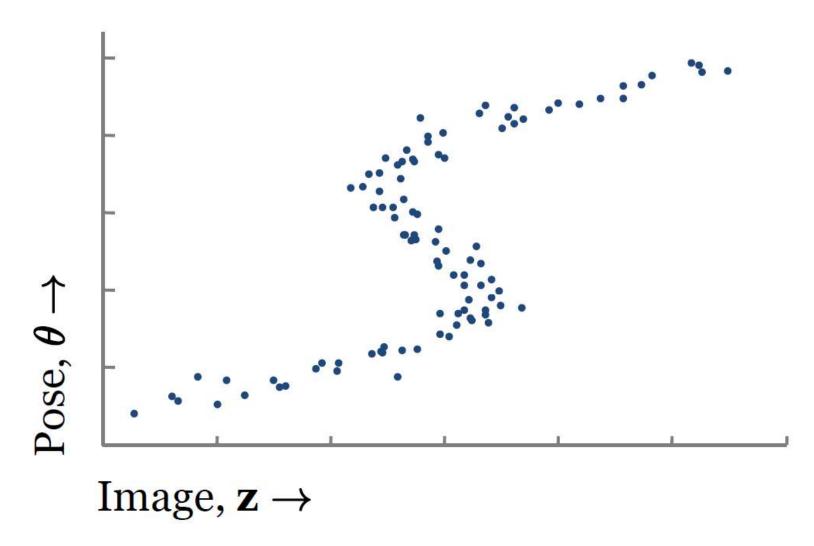


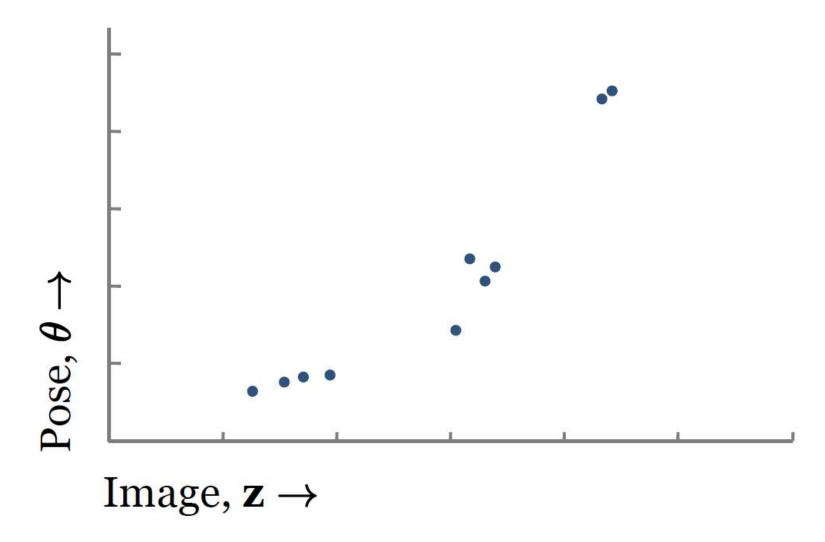
#### For a video sequence:

- Compute modes of conditional at every frame
- Choose sequence of modes to maximize product of likelihood and temporal smoothness using Viterbi

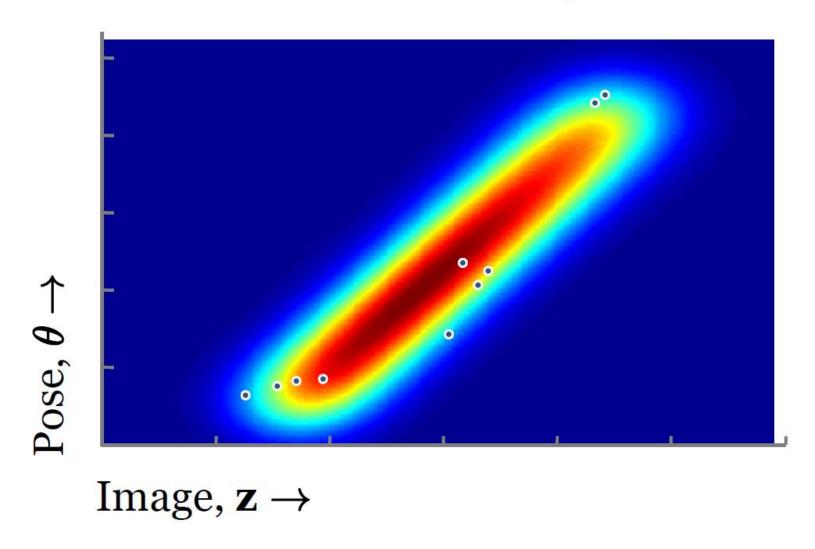
But...

#### Instead of this:

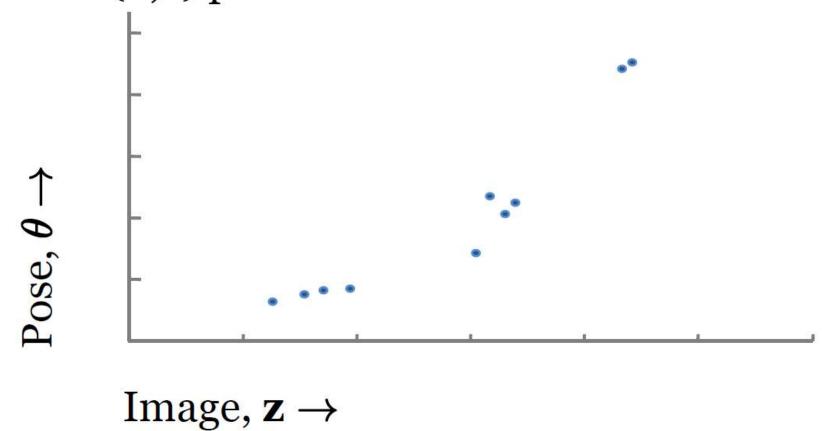




#### Of which a not unreasonable density estimate is:



We have too little training data, i.e. too few labelled ( $\mathbf{z}, \boldsymbol{\theta}$ ) pairs

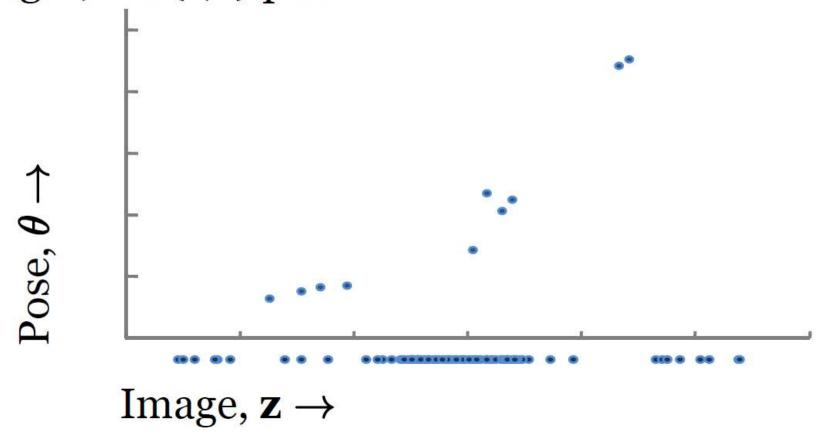


We can't get more because labelling images is expensive...

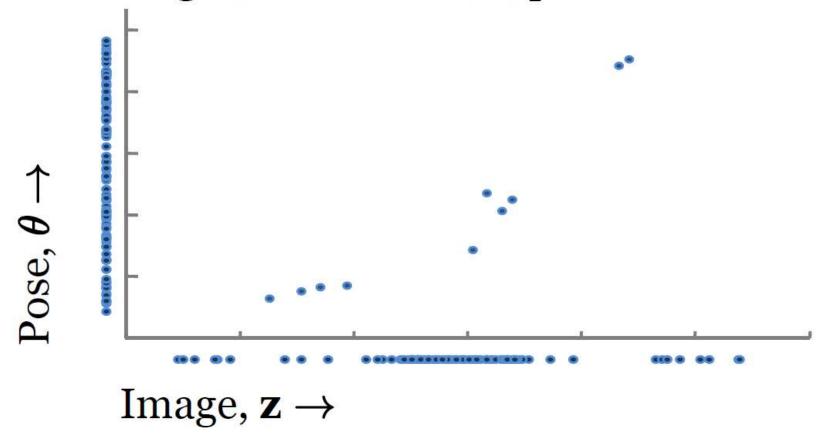
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But we can easily capture more *unlabelled* images, i.e. (z,\*) pairs



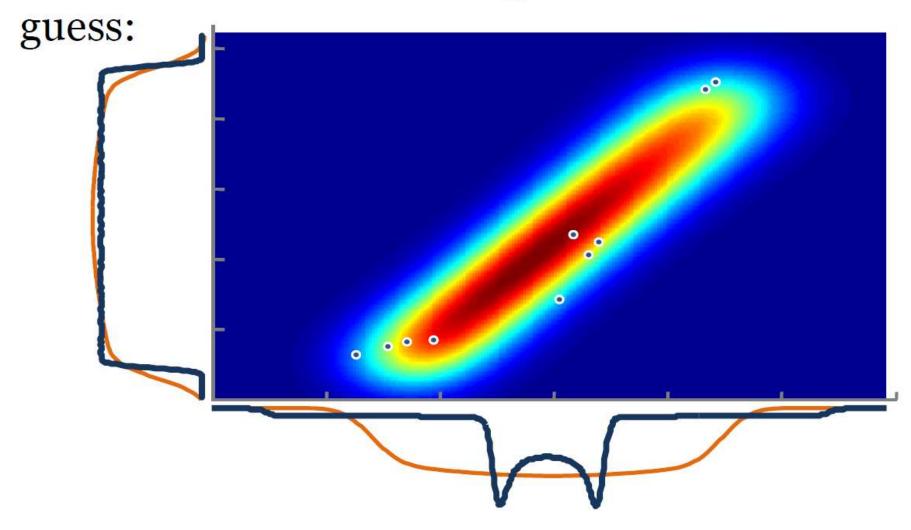
And we can easily download more mocap data without images, i.e. more  $(*,\theta)$  pairs



In fact, it's as if we know the **marginals**  $p(\boldsymbol{\theta}) = \int p(\mathbf{z}, \boldsymbol{\theta}) d\mathbf{z}$  and  $p(\mathbf{z}) = \int p(\mathbf{z}, \boldsymbol{\theta}) d\boldsymbol{\theta}$ Pose,  $p(\theta)$ Image,  $p(\mathbf{z})$ 

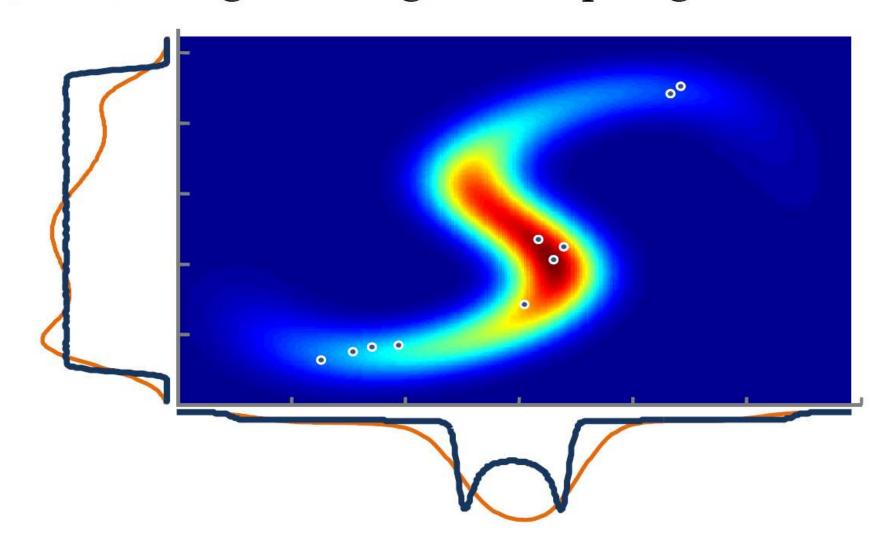


Which contradict the marginals of our earlier





## [ffwd] Using the marginal samples gives this:



#### Joint manifold model

t: a continuous latent variable

t=0

Joint density  $p(\mathbf{z}, \boldsymbol{\theta}) = \int p(\boldsymbol{\theta}|\mathbf{t})p(\mathbf{z}|\mathbf{t})p(\mathbf{t})d\mathbf{t}$ 

Or, loosely, the "spine" of the joint density is a manifold

$$\begin{pmatrix} \boldsymbol{\theta}(\mathbf{t}) \\ \mathbf{z}(\mathbf{t}) \end{pmatrix} \qquad \mathbf{t} \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$$

t=1.0

plus noise.

$$p(z,\theta) = \int p(z,\theta|t)p(t)dt$$

t=-.5

$$t = -1.0$$

#### Joint manifold model

Find latent variables  $\mathbf{t}$  to maximize the posterior of training data  $\mathbf{D} = \{(\boldsymbol{\theta}_l, \mathbf{z}_l)\}_{l=1}^L$ 

$$p(\mathbf{t}_{1..L}|\mathbf{D}) \propto p(\mathbf{D}|\mathbf{t}_{1..L})p(\mathbf{t}_{1..L})$$

$$= p(\boldsymbol{\theta}_{1..L}, \mathbf{z}_{1..L}|\mathbf{t}_{1..L})p(\mathbf{t}_{1..L})$$

$$= p(\boldsymbol{\theta}_{1..L}|\mathbf{t}_{1..L})p(\mathbf{z}_{1..L}|\mathbf{t}_{1..L})p(\mathbf{t}_{1..L}).$$

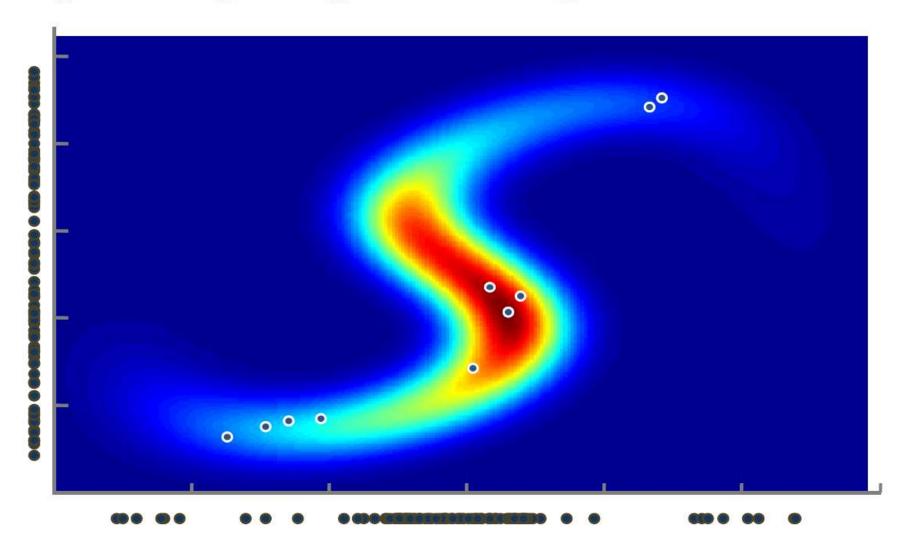
### Adding the marginal samples...

#### Maximize the posterior of

- joint training data  $\mathbf{D} = \{(\boldsymbol{\theta}_l, \mathbf{z}_l)\}_{l=1}^L$
- marginal  $\boldsymbol{\theta}$  data  $\mathbf{M} = \{(\boldsymbol{\theta}_{l}^{*}, *)\}_{l=1}^{L}$
- marginal **z** data  $\mathbf{N} = \{(*, \mathbf{z}_l^*)\}_{l=1}^L$

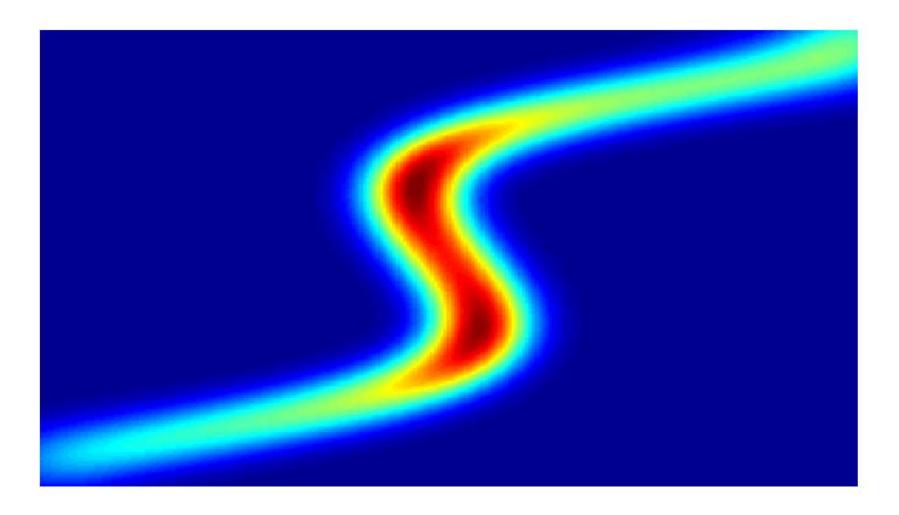
$$p(\boldsymbol{\theta}_{1..L}|\mathbf{t}_{1..L})p(\mathbf{z}_{1..L}|\mathbf{t}_{1..L}) \times \\ p(\boldsymbol{\theta}_{1..L}^*|\mathbf{t}_{1..L}^{\boldsymbol{\theta}}) \times p(\mathbf{z}_{1..L}^*|\mathbf{t}_{1..L}^{\mathbf{z}}) \times \\ p(\mathbf{t}_{1..L}, \mathbf{t}_{1..L}^{\boldsymbol{\theta}}, \mathbf{t}_{1..L}^{\mathbf{z}})$$

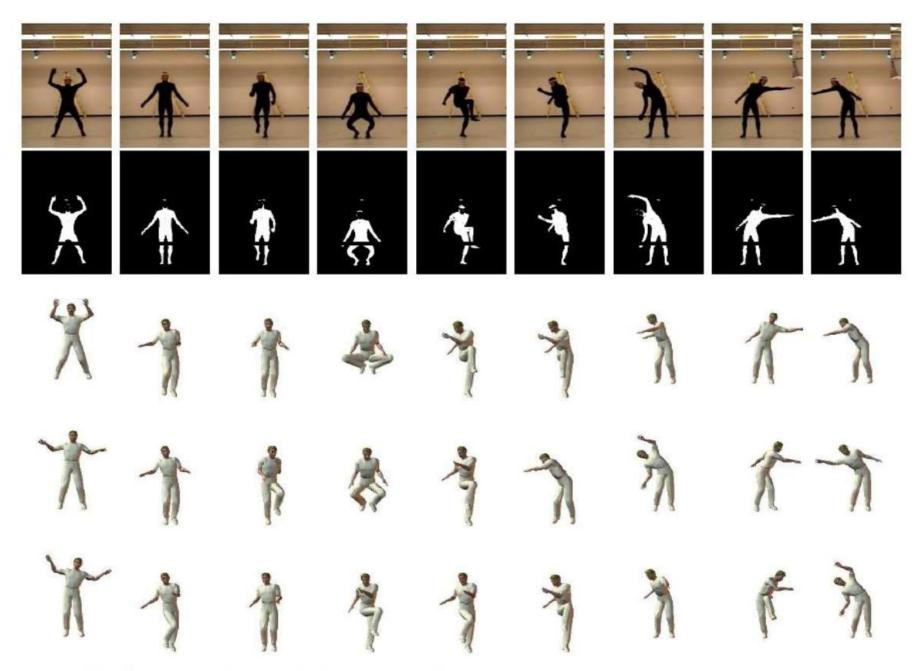
# Optimizing using scaled CG gives:



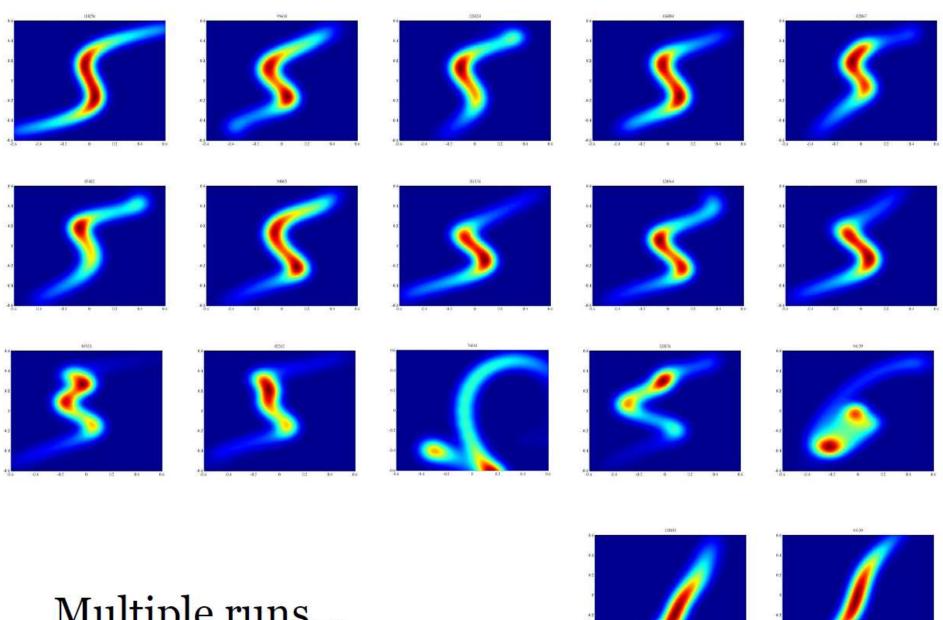


# The estimate from 100 training pairs:



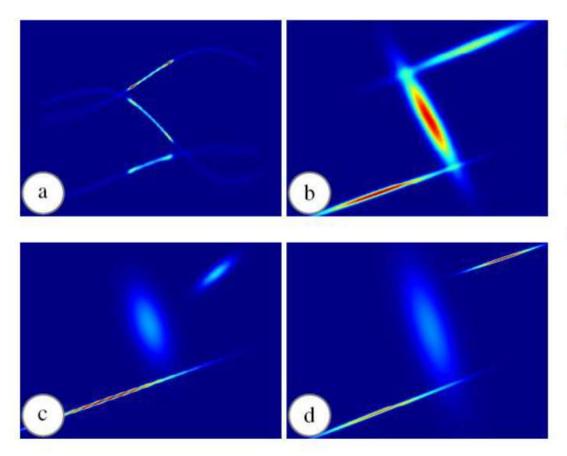


Applied to real-world example



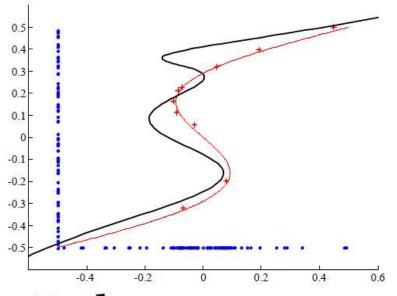
Multiple runs...

#### Other methods

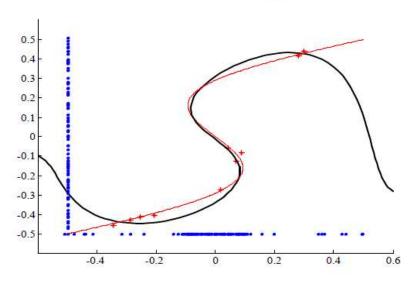


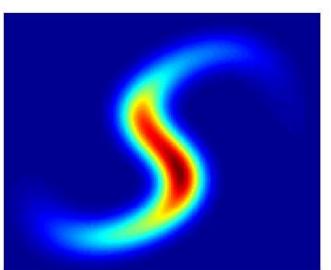
- a. Mixture of experts, 15J
- b. GMM, 15J
- c. GMM, 8J
- d. GMM, 8J, 104M

#### Failure modes



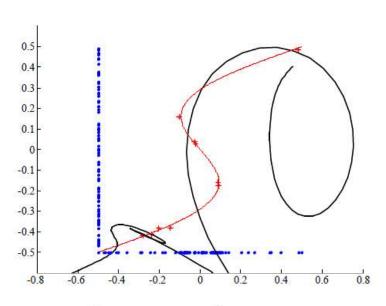
Bad convergence

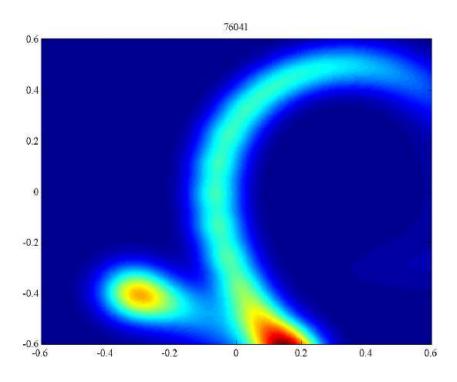




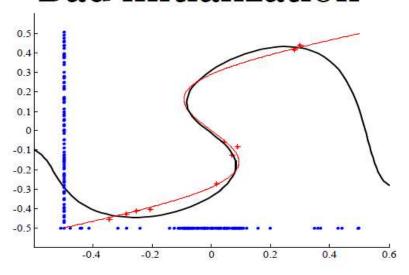
EE4-62 MLCV

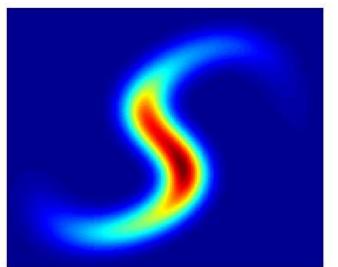
#### Failure modes





#### Bad initialization





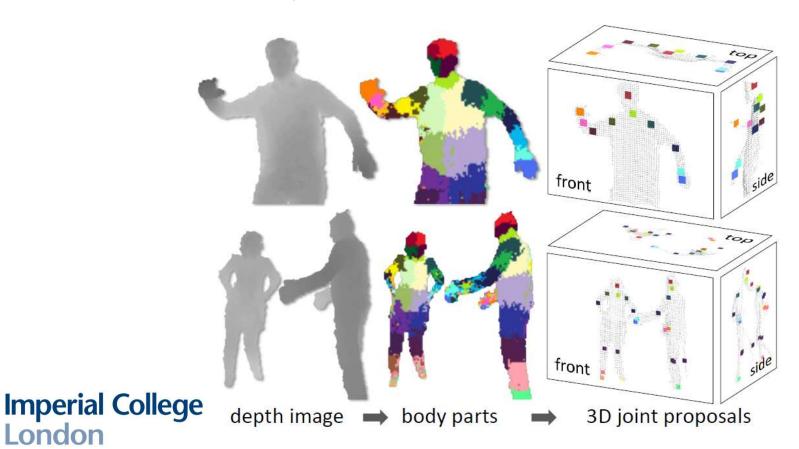
EE4-62 MLCV

### Real-Time Human Pose Recognition in Parts from Single Depth Images (J. Shotton et al, 2011)

Key features

London

- Depth image as input
- Real-time by Random Forest, and Part-based



EE4-62 MLCV

# Progressive Search Space Reduction for Human Pose Estimation (Ferrari et al, 2008)

